

# SDS 384 11: Theoretical Statistics Lecture 6: Lipschitz continuous functions

Purnamrita Sarkar Department of Statistics and Data Science The University of Texas at Austin

## **Recall-Lipschitz functions of Gaussian random variables**

### Definition

A function  $F : \mathbb{R}^n \to \mathbb{R}$  is *L*-Lipschitz w.r.t the Euclidean norm if

$$|F(x) - F(y)| \le L ||x - y||_2 \qquad \forall x, y \in \mathbb{R}^n$$

#### Theorem (LG:Lipschtiz functions of Gaussians)

Let  $(X_1, ..., X_n)$  be a vector of iid N(0, 1) random variables. Let  $f : \mathbb{R}^n \to \mathbb{R}$  be L-Lipschitz w.r.t the Euclidean norm. Then F(X) - E[F(X)] is sub-gaussian with parameter at most L, i.e.  $\forall t \ge 0$ ,

$$P(|F(X) - E[F(X)]| \ge t) \le e^{-\frac{t^2}{2L^2}}$$

• So a *L*-Lipschitz function of *n* gaussian random variables behave like a subgaussian with variance proxy *L*<sup>2</sup>.

# Proof – (Courtesy Tao, Maurey and Pisier)

We will assume that f is differentiable everywhere. So, L Lipschitz simply means that the gradient norm is bounded by L. Part of the reason for this is a consequence of Rademacher's theorem, i.e. Lipschitz continuous functions are differentiable *almost everywhere*.

### Proof.

- WLOG assume E[F(X)] = 0 and L = 1. Assume for simplicity that F is smooth
- We will just prove the upper tail  $P(F(X) \ge \lambda) \le C \exp(-c\lambda^2)$ .
- All we need is

$$E[e^{tF(X)}] \le e^{C't^2} \qquad \text{for } t > 0 \tag{1}$$

• Lipschitz property implies the gradient  $|\nabla F(x)| \leq 1 \forall x \in \mathbb{R}^n$ 

### Proof contd.

### Proof contd.

- Consider an iid copy Y.
- Jensen's inequality implies  $E[e^{-tF(Y)}] \ge e^{-tE[F(Y)]} = 1$

• 
$$E[e^{tF(X)}] \leq E\left[e^{t(F(X)-F(Y))}\right]$$
  
 $F(X) - F(Y) = \int_{0}^{\pi/2} \frac{d}{d\theta} F(\underbrace{X \sin \theta + Y \cos \theta}_{X_{\theta}}) d\theta$   
 $= \frac{\pi}{2} E_{\theta} \left[F'(X_{\theta}) \cdot X'_{\theta}\right]$   
 $e^{t(F(X)-F(Y))} \leq E_{\theta} \left[e^{\frac{\pi}{2}tF'(X_{\theta}) \cdot X'_{\theta}}\right]$ 

•  $X'_{\theta} = X \cos \theta - Y \sin \theta$ . Also note that  $X_{\theta}, X'_{\theta} \stackrel{iid}{\sim} N(0, I_n)$ 

### Proof contd.

### Proof contd.

• 
$$e^{t(F(X)-F(Y))} \leq \frac{2}{\pi} \int_0^{\pi/2} e^{\frac{\pi}{2}tF'(X_{\theta})\cdot X'_{\theta}} d\theta$$
  
•  $X'_{\theta} = X \cos \theta - Y \sin \theta$ . Also note that  $X_{\theta}, X'_{\theta} \stackrel{iid}{\sim} N(0, I_n)$   
 $E[e^{t(F(X)-F(Y))}] \leq \frac{2}{\pi} \int_0^{\pi/2} E[e^{\frac{\pi}{2}tF'(X_{\theta})\cdot X'_{\theta}}] d\theta$   
•  $= \frac{2}{\pi} \int_0^{\pi/2} E_{X_{\theta}} E_{X'_{\theta}} [e^{\frac{\pi}{2}tF'(X_{\theta})\cdot X'_{\theta}}|X_{\theta}] d\theta$   
 $\leq e^{\frac{\pi^2 t^2}{8}}$ 

- The last step is true because conditioned on  $X_{\theta}$ ,  $F'(X_{\theta}) \cdot X'_{\theta} \sim N(0, \sigma^2)$  where  $\sigma \leq 1$ .
- This proves Eq 1.

- Remember our friend chi square r.v.s? Consider  $\{X_i\}_{i=1}^n \stackrel{\text{iid}}{\sim} N(0,1)$ .
- We proved that  $Y = \sum_{i} X_{i}^{2}$  is subexponential and we got the bound  $P(|Y/n - 1| \ge \epsilon) \le 2e^{-n\epsilon^{2}/8}.$
- Lets try to prove a similar bound with the LG theorem.

• Let 
$$\underline{x} = (x_1, \dots, x_n)$$
 and  $f(\underline{x}) = ||\underline{x}||_2$ .

- Note that Euclidian norm is 1-Lipschitz.
- So we have  $P(f(X) E[f(X)] \ge t) \le e^{-t^2/2}$  for  $t \ge 0$ .
- Since  $E[\sqrt{V}] \le \sqrt{E[V]}$ , we have  $E[\sqrt{Y}] \le \sqrt{E[Y]} = \sqrt{n}$ .
- $P(f(X) \ge E[f(X)] + t) \ge P(\sqrt{Y} \ge \sqrt{n} + t) = P(Y/n \ge (1 + \epsilon)^2)$

• Since 
$$(1 + \epsilon/3)^2 \le 1 + \epsilon$$
, for  $\epsilon \in (0, 1)$ ,  
 $e^{-n\epsilon_0^2/18} \ge P(Y/n \ge (1 + \epsilon_0/3)^2) \ge P(Y/n \ge 1 + \epsilon_0)$ 

Consider a sequence of independent standard normal r.v.s  $X = \{X_1, \dots, X_n\}. \text{ Let } X_{(1)} \ge X_{(2)} \ge \dots \ge X_{(n)}.$   $P(|X_{(k)} - E[X_{(k)}]| \ge \epsilon) \le 2e^{-\epsilon^2/2}$ 

#### Proof.

- First note that  $|X_{(k)} Y_{(k)}| \le ||X Y||_2$ . (How?)
- So the order statistics are 1-Lipschitz.

Consider a iid sequence  $X = \{X_i\}_{i=1}^n$ . We will bound  $f(X) := \sup_{a \in \mathcal{A}} a^T X$ where  $\mathcal{A}$  is a compact subset of  $\mathbb{R}^n$  such that  $\mathcal{W} = \sup_{a \in \mathcal{A}} ||a||_2 < \infty$ .

- Why cant we just use Chernoff?
- First let us check if f(X) is Lipschitz. Let  $a_*$  and  $a'_*$  be the maximizers of f(X) and f(X').  $f(X) - f(X') = a^T_* X - a'^T_* X' \le a^T_* (X - X')$  $\le \sup_{a \in \mathcal{A}} a^T (X - X') \le \mathcal{W} ||X - X'||_2$

Consider a iid sequence  $X = \{X_i\}_{i=1}^n$ . We will bound  $f(X) := \sup_{a \in \mathcal{A}} a^T X$ where  $\mathcal{A}$  is a compact subset of  $\mathbb{R}^n$  such that  $\mathcal{W} = \sup_{a \in \mathcal{A}} ||a||_2 < \infty$ .

- If  $X_i \sim N(0,1)$  using Gaussian+Lipschtz
  - $P(|f(X) E[f(X)]| \ge t) \le 2e^{-\frac{t^2}{2W^2}}$
- How about McDiarmid?