



THE UNIVERSITY OF TEXAS AT AUSTIN

Department of Statistics and Data Sciences

College of Natural Sciences

SDS 321: Introduction to Probability and Statistics

Lecture 17: Continuous random variables: conditional PDF

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Roadmap

- ▶ Two random variables: joint distributions
 - ▶ Joint pdf
 - ▶ Joint pdf to a single pdf: Marginalization
 - ▶ Conditional pdf
 - ▶ Conditioning on an event
 - ▶ Conditioning on a continuous r.v
 - ▶ Total probability rule for continuous r.v's
 - ▶ Bayes theorem for continuous r.v's
 - ▶ Conditional expectation and total expectation theorem
 - ▶ Independence
- ▶ More than two random variables.

Conditional PDFs—conditioning on an event

- ▶ For *discrete* random variables, we looked at marginal PMFs $p_X(X)$, conditional PMFs $p_{X|Y}(x|y)$, and joint PMFs $p_{X,Y}(x,y)$.
- ▶ These corresponded to the probability of an event, $P(A)$, the conditional probability of an event given some other event, $P(A|B)$, and probability of the intersection of two events, $P(A \cap B)$.
- ▶ We've looked at marginal PDFs, $f_X(x)$ and joint PDFs, $f_{X,Y}(x,y)$.
- ▶ These don't directly give us probabilities of events, but we can use them to calculate such probabilities by integration.
- ▶ We can also look at conditional PDFs! These allow us to calculate the probability of events given extra information.

Conditional PDFs

- ▶ Recall, the PDF of a continuous random variable X is the non-negative function $f_X(x)$ that satisfies

$$P(X \in B) = \int_B f_X(x) dx$$

for any subset B of the real line.

- ▶ Let A be some event with $P(A) > 0$
- ▶ The **conditional PDF** of X , given A , is the non-negative function $f_{X|A}$ that satisfies

$$P(X \in B | X \in A) = \int_B f_{X|A}(x) dx$$

for any subset B of the real line.

- ▶ If B is the entire line, then we have

$$\int_{-\infty}^{\infty} f_{X|A}(x) dx = 1$$

- ▶ So, $f_{X|A}(x)$ is a valid PDF.

Conditional PDFs

- ▶ The event we are conditioning on can also correspond to a range of values of our continuous random variable.

- ▶ **Definition-**

$$f_{X|\{X \in A\}}(x) = \begin{cases} \frac{f_X(x)}{P(X \in A)} & \text{if } X \in A \\ 0 & \text{otherwise.} \end{cases}$$

- ▶ In this case, we can write the conditional probability as

$$\begin{aligned} \int_B f_{X|A}(x) dx &= \int_B \frac{f_X(x) \mathbf{1}(x \in A)}{P(X \in A)} dx \\ &= \frac{\int_{A \cap B} f_X(x) dx}{P(X \in A)} = \frac{P(\{X \in A\} \cap \{X \in B\})}{P(X \in A)} \\ &= P(X \in B | X \in A) \end{aligned}$$

- ▶ This is a valid PDF—non-negative and integrates to one. Check?

Conditioning: memoryless property of the exponential

- ▶ $X \sim \text{Exp}(\lambda)$
- ▶ $f_X(x) = \lambda e^{-\lambda x}$ when $x \geq 0$, and zero otherwise.
- ▶ $P(X > s + t | X > s) = ?$

Conditioning: memoryless property of the exponential

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- ▶ $f_X(x) = \lambda e^{-\lambda x}$ when $x \geq 0$, and zero otherwise.
- ▶ $P(X > s + t | X > s) = ?$
- ▶ Remember the exponential? $F_X(x) = 1 - e^{-\lambda x}$.

$$P(X > s + t | X > s) = \frac{P(X > s + t, X > s)}{P(X > s)}$$

- ▶
$$= \frac{P(X > s + t)}{P(X > s)} = \frac{e^{-\lambda(s+t)}}{e^{-\lambda s}}$$
$$= e^{-\lambda t} = P(X > t)$$

Conditioning: memoryless property of the exponential

▶ $X \sim \text{Exp}(\lambda)$

▶ $f_{X|X>s}(x) = \begin{cases} \frac{\lambda e^{-\lambda x}}{P(X > s)} = \lambda e^{\lambda(x-s)} & \text{If } x > s \\ 0 & \text{Otherwise} \end{cases}$

▶ $P(X > s + t | X > s) = ?$

Conditioning: memoryless property of the exponential

▶ $X \sim \text{Exp}(\lambda)$

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$$f_{X|X>s}(x) = \begin{cases} \frac{\lambda e^{-\lambda x}}{P(X > s)} = \lambda e^{\lambda(x-s)} & \text{if } x > s \\ 0 & \text{Otherwise} \end{cases}$$

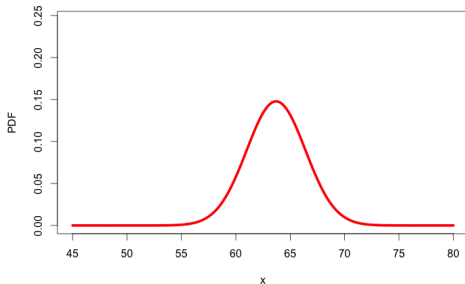
▶ $P(X > s + t | X > s) = ?$

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▶
$$\begin{aligned} P(X > s + t | X > s) &= \int_{s+t}^{\infty} f_{X|X>s}(x) dx = \lambda \int_{s+t}^{\infty} e^{-\lambda(x-s)} dx \\ &= \lambda \int_t^{\infty} e^{-\lambda u} du = e^{-\lambda t} \end{aligned}$$

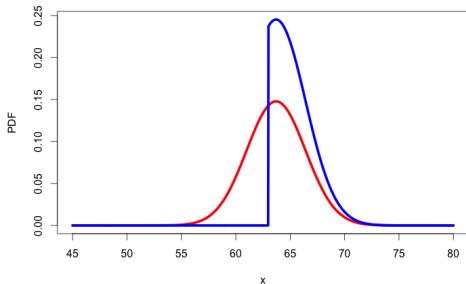
Conditional PDFs: Example

- ▶ The height X of a randomly picked american woman can be modeled by $X \sim N(63.7, 2.7^2)$
- ▶ Whats the conditional PDF given that the randomly picked woman is at least 63 inches tall?
- ▶ The PDF of heights (X) is shown in red.



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- ▶ Whats the conditional PDF given that the randomly picked woman is at least 63 inches tall?
- ▶ The PDF of heights (X) is shown in red.
- ▶ The conditional PDF given $X > 63$, shown in blue, is the same shape for $X > 63$... but scaled up to integrate to one.



Conditioning on a different random variable

- ▶ So far, we conditioned X on an arbitrary event A , or on a range of values of X .

$$P(X \in B|A) = \int_B f_{X|A}(x)dx$$

- ▶ We can also condition on the outcome of a second random variable Y .
- ▶ We know we could condition on a range of outcomes of Y , by replacing the arbitrary event A with the event $\{Y \in A\}$

$$P(X \in B|Y \in A) = \int_B f_{X|\{Y \in A\}}(x)dx$$

- ▶ What about conditioning on a specific value of $Y = y$?
- ▶ Even though any outcome $Y = y$ has $P(Y = y) = 0$, we know that *some* value has to happen.
 - ▶ Pick some number, say 0.6777, now generate 100 $N(0, 1)$ random variables. I will bet a 100\$ that you won't see that number.
 - ▶ But when you simulate from the standard normal, you will get a 100 different values, right?

Conditioning on a different random variable



$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)},$$

provided $f_Y(y) > 0$.

- ▶ What does this mean?

$$f_{X|Y}(x|y)dx = \frac{f(x,y)dxdy}{f(y)dy}$$



$$\begin{aligned} &= \frac{P(x \leq X \leq x + dx, y \leq Y \leq y + dy)}{P(y \leq Y \leq y + dy)} \\ &= P(x \leq X \leq x + dx | y \leq Y \leq y + dy) \end{aligned}$$

Multiplication rule: Calculating the joint PDF

- ▶ We can use the same relationship, $f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$, to calculate the joint PDF from the conditional and the marginal PDF.
- ▶ i.e., $f_{X,Y}(x,y) = f_{X|Y}(x|y)f_Y(y)$.
- ▶ This is a PDF version of our multiplication rule.
- ▶ We can extend it to more than 2 random variables:

$$f_{X,Y,Z}(x,y,z) = f_{Z|X,Y}(z|x,y)f_{Y|X}(y|x)f_X(x)$$

Lets remember all the rules

We've now got a lot of ways to go between our various PDFs!

- ▶ If we know $f_{X,Y}(x,y)$, we can get $f_X(x)$
 - ▶ How?
- ▶ If we know $f_{X,Y}(x,y)$ and $f_Y(y)$, if $f_Y(y) > 0$ we can get $f_{X|Y}(x|y)$
- ▶ If we know $f_X(x)$ and $f_{Y|X}(y|x)$, we can get $f_{X,Y}(x,y)$

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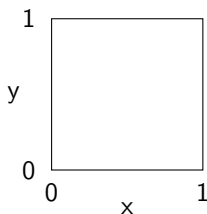
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- ▶ If we know $f_X(x)$ and $f_{Y|X}(y|x)$, we can get $f_{X,Y}(x,y)$
 - ▶ How? multiplication rule! $f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x)$

Example: Calculating the conditional PDF

- ▶ Let $f_{X,Y}(x,y) = \begin{cases} c & \text{if } 0 \leq x \leq 1 \text{ and } 0 \leq y \leq x \\ 0 & \text{otherwise} \end{cases}$
- ▶ What is the conditional PDF of X given Y , $f_{X|Y}(x|y)$?

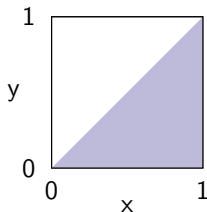
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- ▶ First things first... what is c ? Well, what does our joint PDF look like?



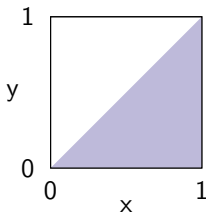
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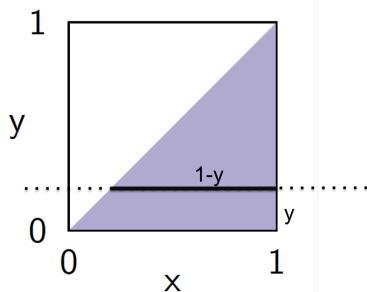
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- ▶ The total area where $0 \leq x \leq 1$ and $0 \leq y \leq x$ is 0.5, so $c = 2$.
- ▶ What is the marginal PDF of Y , $f_Y(y)$?

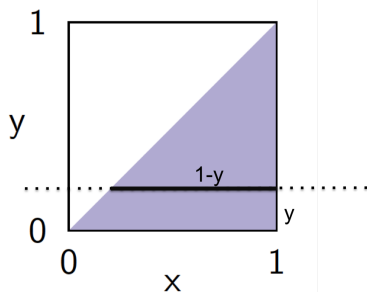
Example: Calculating the conditional PDF



- ▶ To get the marginal PDF of Y , we take the joint PDF and marginalize out X .

- ▶
$$f_Y(y) = \int_0^1 f_{X,Y}(x,y) dx = 2 \int_0^1 \mathbf{1}_{0 \leq x \leq 1, 0 \leq y \leq x} dx$$

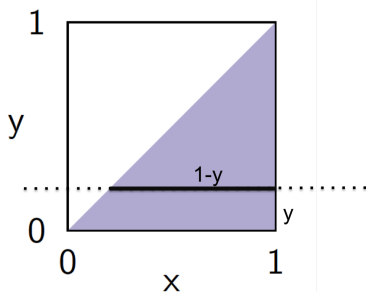
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$$= 2 \int_{x=y}^1 dx = 2(1-y)$$

Example: Calculating the conditional PDF



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$$= 2 \int_{x=y}^1 dx = 2(1-y)$$

- ▶ So, the conditional PDF of X given $Y = y$ is

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \begin{cases} \frac{1}{1-y} & \text{if } y \leq x \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

Total probability theorem for continuous random variables

- ▶ We know that conditional probabilities must obey the total probability theorem.
- ▶ If B_1, \dots, B_n form a partition of Ω , such that $P(B_i) > 0$ for each i , then for any event A ,

$$P(A) = \sum_{i=1}^n P(B_i)P(A|B_i)$$

- ▶ In terms of discrete r.v.'s we have:

$$P(X = x) = \sum_i P(X = x|B_i)P(B_i)$$

- ▶ How about continuous r.v.'s? Replace $P(X = x|B_i)$ by conditional pdf.

$$f_X(x) = \sum_i f_{X|B_i}(x)P(B_i)$$

Bayes' law with continuous outcomes but discrete hidden causes

- ▶ Sometimes our hidden cause is inherently discrete.
 - ▶ e.g. I may be interested in whether I have flu or not – a binary choice.
 - ▶ My observation might be my temperature – a continuous random variable.
- ▶ We want $P(A|Y = y) =$ e.g. $P(\text{flu}|Y = 100)$
- ▶ Pretend Y is a discrete r.v.

$$P(A|Y = y) = \frac{P(Y = y|A)P(A)}{P(Y = y|A)P(A) + P(Y = y|A^c)P(A^c)}$$

All that changes for a continuous r.v. is:

$$P(A|Y = y) = \frac{f_{Y|A}(y)P(A)}{f_{Y|A}(y)P(A) + f_{Y|A^c}(y)P(A^c)}$$

Bayes' law with continuous outcomes but discrete hidden causes

- ▶ The probability that anyone has flu (event A) is 20%.
- ▶ Body temperature is Y .
- ▶ Without flu, Y is a normal random variable with $\mu = 98.6$ degrees and $\sigma = .5$.
- ▶ With flu, Y is a normal random variable with $\mu = 102$ and $\sigma = 2$.
- ▶ My temperature is 100. If A is the event "has flu" and Y is temp.

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$$f_{Y|A}(y) = \frac{1}{\sqrt{2\pi \times 4}} \exp - \frac{(y - 102)^2}{2 \times 4}$$

$$f_{Y|A^c}(y) = \frac{1}{\sqrt{2\pi \times .25}} \exp - \frac{(y - 98.6)^2}{2 \times .25}$$

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$$f_{Y|A^c}(y) = \frac{1}{\sqrt{2\pi \times .25}} \exp - \frac{(y - 98.6)^2}{2 \times .25}$$

$$P(A|Y = y) = \frac{P(A)f_{Y|A}(y)}{f_Y(y)} = \frac{f_{Y|A}(y)P(A)}{f_{Y|A}(y)P(A) + f_{Y|A^c}(y)P(A^c)}$$
$$P(A|Y = 100) = \frac{0.2 \frac{1}{2\sqrt{2\pi}} e^{-(100-102)^2/8}}{0.2 \frac{1}{2\sqrt{2\pi}} e^{-(100-102)^2/8} + 0.8 \frac{1}{0.5\sqrt{2\pi}} e^{-(100-98.6)^2/0.5}} = 0.65$$

Continuous Bayes' rule

- ▶ Discrete X, Y .

- ▶
$$P(X = x|Y = y) = \frac{P(Y = y|X = x)P(X = x)}{\sum_x P(Y = y|X = x)P(X = x)}$$

- ▶ What is $f_{X|Y}(x|y)$?

Continuous Bayes' rule

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- ▶
$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{\int_{-\infty}^{\infty} f_{Y|X}(y|x)f_X(x)dx}$$

Conditional Expectation

- ▶ When we were looking at discrete random variables, we looked at **conditional expectations**.
- ▶ The conditional expectation, $E[X|A]$, of a random variable X given an event A is the value of X we expect to get out, on average, when A is true.
- ▶ We could calculate it by summing over all values x that X can take on, and scaling them by the conditional PMF $p_{X|A}(x) = P(X = x|A)$.

$$E[X|A] = \sum_x x p_{X|A}(x)$$

Conditional Expectation

- ▶ We can also look at the conditional expectation of a continuous random variable.
- ▶ If $E[X] = \int_{-\infty}^{\infty} xf_X(x)dx$, what do you think the conditional expectation of X given some event A looks like?

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- ▶ $E[X|A] = \int_{-\infty}^{\infty} xf_{X|A}(x)dx$

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- ▶ $E[X|A] = \int_{-\infty}^{\infty} xf_{X|A}(x)dx$
- ▶ How about the conditional expectation of some function $g(X)$ given some event A ?

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- ▶ $E[X|A] = \int_{-\infty}^{\infty} xf_{X|A}(x)dx$
- ▶ How about the conditional expectation of some function $g(X)$ given some event A ?
- ▶ $E[g(X)|A] = \int_{-\infty}^{\infty} g(x)f_{X|A}(x)dx$

Total expectation theorem

- ▶ More generally, if A_1, A_2, \dots, A_n are a partition of Ω , we have a continuous version of the **total expectation theorem**:

$$E[X] = \sum_{i=1}^n P(A_i)E[X|A_i]$$

- ▶ Or, if we are conditioning on specific values $Y = y$,

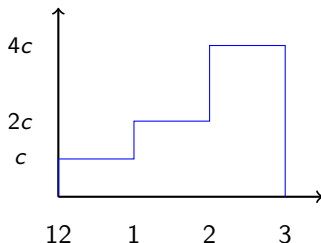
$$E[X] = \int_{-\infty}^{\infty} E[X|Y = y]f_Y(y)dy$$

Conditional expectation

- ▶ I am expecting an email, that will definitely arrive between midday and 3pm.
- ▶ Within a given hour (midday-1, 1-2, 2-3), each time is equally likely.
- ▶ It is twice as likely to arrive between 1 and 2 as it is to arrive between midday and 1.
- ▶ It is twice as likely to arrive between 2 and 3 as it is to arrive between 1 and 2.
- ▶ What does the PDF look like?

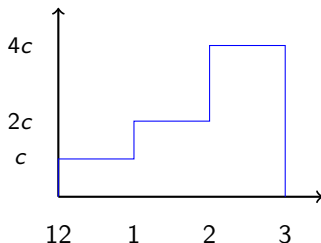
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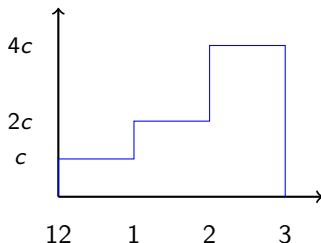
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- ▶ What is c ?

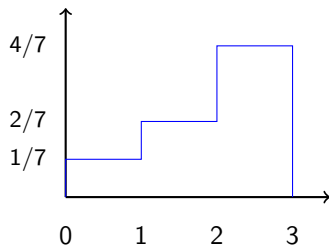
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- ▶ It is twice as likely to arrive between 2 and 3 as it is to arrive between 1 and 2.
- ▶ What does the PDF look like?



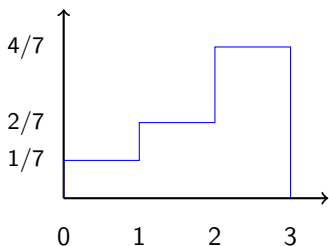
- ▶ What is c ? $1/7$

Conditional expectation



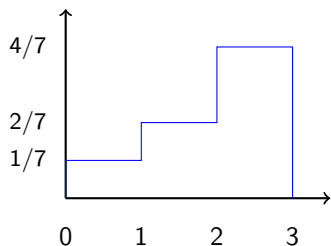
- ▶ I wait until 2pm. It still hasn't arrived. What is the expected value of the arrival time?
- ▶ What is the expected time without any conditioning?

Conditional expectation



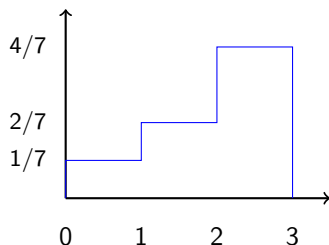
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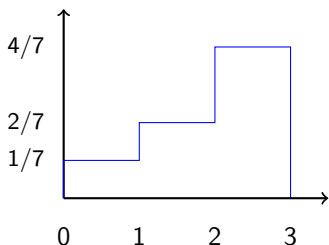
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$$f_{X|X>2}(x) = \begin{cases} 1 & \text{if } 2 < x \leq 3 \\ 0 & \text{otherwise} \end{cases}$$

Conditional expectation

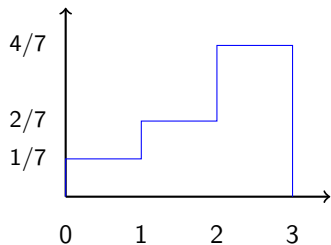


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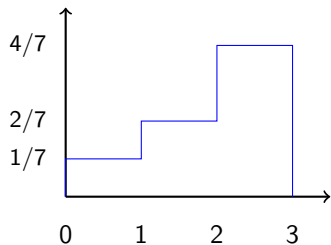
- ▶ So, $E[X|X > 2] = \int_{-\infty}^{\infty} x f_{X|X>2}(x) dx = \int_2^3 x dx = 2.5$.

Conditional expectation



- ▶ What is the (unconditional) probability that $X > 2$?

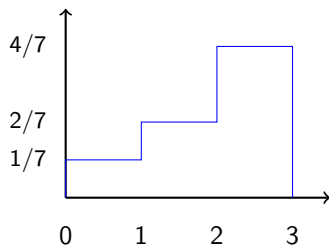
Conditional expectation



- ▶ What is the (unconditional) probability that $X > 2$?

- ▶
$$P(X > 2) = \int_2^3 f_X(x) dx = 4/7$$

Conditional expectation



- ▶ What is the (unconditional) probability that $X > 2$?
- ▶ $P(X > 2) = \int_2^3 f_X(x) dx = 4/7$
- ▶ Similarly, $P(X < 1) = \int_0^1 f_X(x) dx = 1/7$ and $P(1 \leq X \leq 2) = 2/7$.

Total expectation theorem

- ▶ What is the total expectation of X ?

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Total expectation theorem

- ▶ What is the total expectation of X ?
- ▶ $E[X] = \int_{-\infty}^{\infty} xf_X(x)dx$
- ▶ By the total probability theorem,

$$f_X(x) = P(X \leq 1)f_{X|0 \leq X \leq 1}(x) \\ + P(1 \leq X \leq 2)f_{X|1 \leq X \leq 2}(x) + P(X > 2)f_{X|X > 2}(x)$$

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- ▶ So, we can write the total expectation as

$$E[X] = \int_0^1 xP(X \leq 1)f_{X|X \leq 1}(x) + \int_1^2 xP(1 \leq X \leq 2)f_{X|1 \leq X \leq 2}(x) \\ + \int_2^3 xP(X > 2)f_{X|X > 2}(x) \\ = E[X|X \leq 1]P(X \leq 1) + E[X|1 \leq X \leq 2]P(1 \leq X \leq 2) \\ + E[X|X > 2]P(X > 2) \\ = 0.5 \cdot 1/7 + 1.5 \cdot 2/7 + 2.5 \cdot 4/7 = 27/14$$

Total expectation theorem: Example

- ▶ John's tank holds 15 gallons of gas, and he always refills his tank when he gets down to 5 gallons.
- ▶ John's car gets 30MPG on average, with a standard deviation of 2MPG.
- ▶ I plan on borrowing John's car tomorrow. I don't know how much gas he will have. How far should I expect to be able to drive it?
- ▶ Let's set up some reasonable modeling assumptions.
- ▶ Let G be the random volume of gas. Assume

$$f_G(g) = \begin{cases} 0.1 & \text{if } 5 < g \leq 15 \\ 0 & \text{otherwise.} \end{cases}$$

- ▶ I want $E[M]$.
- ▶ Let M be the random number of miles. Assume $M \sim N(30g, 4)$.

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- ▶ If we have exactly g gallons, what is $E[M|G = g]$? **30g**
- ▶ So, we can use the total expectation theorem to get:

$$E[M] = \int_{-\infty}^{\infty} E[M|G = g]f_G(g)dg = \int_5^{15} 30g \times 0.1 dg = [1.5g^2]_5^{15} = 300$$

Independent random variables

- ▶ For discrete random variables, we said two random variables X and Y are independent if

$$p_{X,Y}(x,y) = p_X(x)p_Y(y) \quad \forall x,y$$

- ▶ Just like in the discrete case, we say two continuous random variables are independent if

$$f_{X,Y}(x,y) = f_X(x)f_Y(y) \quad \forall x,y$$

- ▶ If $f_Y(y) > 0$, this is the same as saying $f_X(x) = f_{X|Y}(x|y)$ – i.e. knowing that $Y = y$ doesn't tell us anything about X .
- ▶ Just like with discrete random variables, we if X and Y are independent we have $E[XY] = E[X]E[Y]$ and $\text{var}(X + Y) = \text{var}(X) + \text{var}(Y)$.
 - ▶ For two functions $f(X)$ and $g(Y)$ we have $E[f(X)g(Y)] = E[f(X)]E[g(Y)]$.

More than two random variables

- ▶ For multiple random variables we have:

$$P((X, Y, Z) \in B) = \int_{(x,y,z) \in B} f_{X,Y,Z}(x, y, z) dx dy dz$$

- ▶ Marginalization: $f_{X,Y}(x, y) =$

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- ▶ Multiplication rule:

$$f_{X,Y,Z}(x, y, z) = f_{X|Y,Z}(x|y, z) f_{Y|Z}(y|z) f_Z(z), \text{ For } f_{Y,Z}(y, z) > 0$$

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- ▶ Independence: $f_{X,Y,Z}(x, y, z) = f_X(x) f_Y(y) f_Z(z)$ For all x, y, z

More than two random variables

- ▶ For two random variables X, Y arising out of the same experiment, we define their CDF as:

$$F_{X,Y}(x,y) = P(X \leq x, Y \leq y) =$$

- ▶ How do I get $f_{X,Y}(x,y)$ back? $f_{X,Y}(x,y) = \frac{d^2 F_{X,Y}(x,y)}{dx dy}$
- ▶ Let X and Y be jointly uniform on the unit square. $F_{X,Y}(x,y) = xy$ for $0 \leq x, y \leq 1$
- ▶ What is $f_{X,Y}(x,y)$?. Differentiate! $\frac{d}{dx} \left(\frac{d}{dy}(xy) \right)$
- ▶ This equals 1 for all $0 \leq x, y \leq 1$!

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$$F_{X,Y}(x,y) = P(X \leq x, Y \leq y) = \int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(u,v) du dv$$

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Practice problem

- ▶ Let $Y = g(X) = X^2$. X is a random variable with a known PDF $f_X(x)$. What's the PDF of Y ?
- ▶ Solution: See example 3.23 of Bertsekas and Tsitsiklis.